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2	Trial history contributes to the optimal tuning of attention
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10	Abstract: 150 words
11	Significance statement: 95 words
12	Main text: 6621 words
13	Figures: 4
14	Tables: 2
15	Running head: Optimal tuning and trial history
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23 Abstract In visual search tasks, targets are difficult to find when they are similar to the 24 surrounding nontargets. In this scenario, it is optimal to tune attention to target features 25 that maximize the difference between target and nontargets. We investigated whether the 26 optimal tuning of attention is driven by biases arising from previously attended stimuli (i.e., 27 28 trial history). Consistent with effects of trial history, we found that optimal tuning was stronger when a single target-nontarget relation was repeated than when two target-29 30 nontarget relations alternated randomly. Detailed analysis of blocks with random alternation showed that optimal tuning was stronger when the target-nontarget relation probed on the 31 current trial matched the relation on the previous trial. We evaluated several mechanisms 32 33 that may underlie effects of trial history, such as priming of attentional set, switch costs, and sensory adaptation. However, none of the accounts was able to fully account for the pattern 34 of results. 35 **Keywords** 36 visual search, attentional capture, optimal tuning, sensory adaptation 37 Public Significance Statement 38 39 Visual search is a common everyday activity. For instance, we often look for objects

in a particular color. This task is difficult when the color of the target is similar to the
surrounding colors. Observers perform surprisingly well in this situation, suggesting that they
adjust attention optimally to maximize the difference between the color of the target and
the surrounding colors. However, we show that this adjustment breaks down when
observers attempt to find more than a single target. Therefore, to optimize performance, it
may best to search for only one object at a time.

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--2--

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Introduction

48	Navalpakkam and Itti (2007) illustrated the optimal tuning of attention using an
49	image where a tiger is mostly hidden in high grass. Because of the importance of spotting
50	the predator, attention should be optimally tuned to detect its features. An analysis of the
51	color statistics of the image revealed that the color of the grass ranged from yellow to green,
52	whereas the color of the tiger ranged from yellow to orange. To increase the chances of
53	finding the tiger, it would not be optimal to look for the average yellow-orange color of the
54	tiger's coat, because the grass is also yellow, and the overlap would result in a poor signal to
55	noise ratio. Instead, it would be optimal to look for orange because orange is not contained
56	in the color of the grass. To provide evidence for the optimal tuning of attention,
57	Navalpakkam and Itti (2007) asked their participants to search for a color target among
58	nontargets in a slightly different color. The search trials were interspersed with probe trials,
59	where participants were shown several colors and had to select the color corresponding to
60	the target. Participants did not choose the true target color most frequently, but a color that
61	was shifted away from the nontarget color. Navalpakkam and Itti's results show that the
62	nontarget color in the search task modified the memory representation of the target color.
63	This memory error has been replicated in numerous experiments (Chapman et al., 2023;
64	Hamblin-Frohman & Becker, 2021; Jung et al., 2021; Kerzel, 2020; Maith et al., 2021; Scolari
65	& Serences, 2009; Yu & Geng, 2019).
66	Navalpakkam and Itti (2007) did not comment on the origins of optimal tuning, but

the literature suggests that there are at least three possibilities. Like attentional selection in general, optimal tuning may be goal-driven, stimulus-driven, or history-driven (e.g., Awh et al., 2012; Liesefeld et al., 2024). Because the memory representation of the target was shifted, it may be that optimal tuning is goal-driven because memory representations are

--3--

used to set the search goals of the observer (Carlisle et al., 2011; Duncan & Humphreys,
1989; Eimer, 2014; Huynh Cong & Kerzel, 2021; Schneider, 2013). In contrast, a stimulusdriven origin is unlikely because the shift is not a characteristic of the stimulus, but a change
in the observer's representation of the stimulus. In the current study, however, we explore
contributions from trial history, which have not been considered so far.

76 Mixed vs. fixed target-nontarget relations

77 The potential role of trial history emerges from a series of studies reporting 78 conflicting results. First, Scolari and Serences (2009) measured contrast detection thresholds 79 to masked stimuli. If attentional selection was biased away from the nontargets, contrast thresholds should be lower for stimuli shifted away from the nontargets because of 80 attentional enhancement (Carrasco, 2011). However, there was no behavioral evidence for 81 82 this prediction, even though a later brain imaging study found some support (Scolari et al., 2012). Importantly, Scolari and Serences (2009) randomly mixed the target and nontarget 83 features to avoid sensory adaptation (see p. 11941). Second, Kerzel (2020) measured cueing 84 effects. Cueing effects are largest for cue colors corresponding to the search goals of the 85 observer (Folk & Remington, 1998). Kerzel (2020) found the largest cueing effects for colors 86 87 shifted away from the nontargets, providing evidence for optimal tuning. Third, Hamblin-Frohman and Becker (2021) measured oculomotor capture by distractors shown 88 89 simultaneously with target and nontarget stimuli. Oculomotor capture is strongest for stimuli that correspond to the search goals of the observer (Wu & Remington, 2003). 90 91 Hamblin-Frohman and Becker (2021) found more oculomotor capture by distractors shifted 92 away from the nontargets, again providing support for optimal tuning. Importantly, target 93 and nontarget colors were fixed across trials in the two studies reporting support for optimal 94 tuning. As a result, the target-nontarget relation was repeated throughout the experiment.

In contrast, the study by Scolari and Serences (2009) did not find support for optimal tuning
when the target-nontarget relations were randomly mixed.

97 Thus, it appears that the optimal tuning of attention depends on trial history and only occurs when a fixed target-nontarget relation is repeated, but previous studies did not 98 directly address this question. Applied to the example of the tiger in the grass, we may only 99 100 be able to optimally tune attention to the tiger if we repeatedly look at the grasslands. If we switch between the grasslands with the tiger and a muddy pond with an almost submerged 101 hippo, we may not be able to optimally tune attention to both targets. For the tiger, the 102 optimally tuned target is "more orange" in a yellowish scene. For the hippo, the skin around 103 the hippo's eyes is red-brown and the pond is brown. The optimally tuned target is therefore 104 105 "redder" in a brownish scene. The question is whether the optimally tuned targets are available when we switch between the grasslands and the pond or whether they are only 106 available when we continuously look at one of the two scenes. In other words, does optimal 107 tuning occur with fixed and mixed target-nontarget relations or is it limited to fixed 108 relations? 109

110 The contingent capture paradigm

111 To answer this question, we used the contingent capture paradigm developed by Folk 112 et al. (1992). In the contingent capture paradigm, spatial cues are shown briefly before the 113 target (see Figure 1A). Cues at the target location are referred to as valid, and cues at a nontarget location as invalid. Although the cues do not predict the target location, valid cues 114 result in shorter RTs than invalid cues. However, these cueing effects only occur when the 115 116 cue color corresponds to the target color. In contrast, cueing effects are absent when the color of the cue does not match the color of the target, suggesting that only matching cues 117 capture attention (e.g., Becker et al., 2019; Carmel & Lamy, 2015; Folk & Remington, 1998; 118

Goller et al., 2020; Kim et al., 2019; Rigsby et al., 2023; Ruthruff et al., 2020; Schönhammer
et al., 2020; Zivony & Lamy, 2018).

Previous research showed that effects of trial history on cueing effects depend on the 121 search task. In a relevant study, Folk and Remington (2008) presented a colored target 122 123 among white nontargets (i.e., a color singleton). Because it was not necessary to know the 124 target color and the target color varied unpredictably between green and red, participants 125 may have looked for any color singleton instead of looking for a particular color. Consistent 126 with singleton search, cueing effects were observed regardless of the match between cue and target color. Further, cueing effects were larger when the cue color on the current trial 127 matched the target color on the previous trial. When the task was changed to favor search 128 129 for a particular color, however, these effects of trial history disappeared. Thus, effects of 130 trial history on cueing effects occur when participants search for singletons, but not when they searched for a particular feature. 131

These results are important because participants in previous studies on optimal tuning also searched for singletons, albeit with color differences that were much smaller than in Folk and Remington (2008). Therefore, it appears likely that optimal tuning is also susceptible to trial history. Possibly, optimal tuning is stronger if the cue on the current trial matches the target-nontarget relation on the previous trial.

We used cueing effects as a diagnostic tool to assess optimal tuning. If optimal tuning occurs, larger cueing effects are expected for cue colors that are shifted away from the nontarget color compared with cue colors that are shifted toward the nontarget color (see Figure 1B). Distances between colors are measured in degrees of rotation on an isoluminant color wheel. Based on prior work (Kerzel, 2020), the optimal tuning was estimated to be about -15° away from the nontarget color with these colors. The tuning of attention can be

--6--

quantified by the difference between cueing effects for cue colors shifted away (-15°) and
toward (+15°) the nontarget color, which we refer to as tuning scores (see Figure 1C). The
larger the tuning scores, the stronger the optimal tuning. In contrast, tuning sores of zero
would indicate that optimal tuning is absent, and that attention is tuned to the exact target
feature. Finally, negative tuning scores would indicate tuning towards the nontarget color,
which is the opposite of optimal tuning.

149 Hypotheses

150 Effects of trial history on optimal tuning were evaluated in two different ways. First, we compared fixed and mixed blocks of trials. In fixed blocks, there was only one target-151 nontarget relation that was repeated on every trial. In mixed blocks, there were two target-152 153 nontarget relations in random order, resulting in some trials where the target-nontarget relation was repeated and others where it changed. If optimal tuning was history-based and 154 depended on the match between the cue on the current trial and the target-nontarget 155 relation on the previous trial, stronger optimal tuning should occur in fixed than in mixed 156 blocks. Second, in mixed blocks, we analyzed optimal tuning separately for trials where the 157 target-nontarget relation was repeated and where it changed. If optimal tuning was history-158 159 based, stronger optimal tuning is expected on trials where the target-nontarget relation was repeated because the cue on the current trial matches the target-nontarget relation on the 160 161 previous trial. In our experiments, the cue always matched the target-nontarget relation on the current trial, but similar results are expected for nonmatching cues because participants 162 163 performed singleton search (see above presentation of Folk & Remington, 2008). 164 Our main hypotheses were about effects of trial history on optimal tuning with two

target-nontarget relations. With two target-nontarget relations, several second-order
 relations are possible. We did not have specific hypotheses about these second-order

--7--

relations but distinguished three which we decided to explore (see Figure 2). Either the two target-nontarget relations shared the same nontarget color (Experiment 1), they shared no color at all (Experiment 2), or they shared the same target color (Experiment 3). Each of these second-order relations is illustrated by a real-world scenario in the introduction to the respective experiment. We expect that the manipulation of second-order relation would shed some light on the processes underlying potential effects of trial history.

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Experiment 1

In Experiment 1, the nontarget color was fixed and the target color varied by +/- 30° around this color, thereby creating two target-nontarget relations. In the example from the introduction, this situation corresponds to a search for two different animals in the same yellow-green grasslands. In addition to the tiger, which is more orange than the grass, there may be a green snake, which is greener than the grass. Of note, these target-nontarget relations are opposite. That is, one is more orange, and the other is greener than the same yellow-green grasslands.

181 Effects of trial history should result in larger optimal tuning in fixed blocks with one 182 target-nontarget relation than in mixed blocks with two target-nontarget relations.

183 Moreover, in mixed blocks, larger optimal tuning is expected when the target-nontarget

relation repeats compared to when it changes because the cue on the current trial matches

185 the target-nontarget relation on the previous trial.

186 Method

Transparency and Openness Promotion. The number of participants was partially
 determined by the counterbalancing demands described below. To fulfill these
 requirements, 24 undergraduate psychology students participated (0 male; age: *M* = 20.5
 years, *SD* = 2.3). In a related study by Kerzel (2020), the sample size was between 18 - 21 and

--8--

the η_p^2 for tuning scores was between .428 - .592, which would require only five 191 participants. In the current study, however, we were interested in modulations of tuning 192 193 scores by intertrial transition. We mostly performed the critical comparisons by paired t-test. G*Power 3.1 (Faul et al., 2009) indicated that we could detect effect sizes as small as d_z = 194 0.66 with a sample size of 20 (two-tailed, alpha = .05, power = .8). None of the experiments 195 196 reported in this article was formally preregistered. The data are available at 197 https://osf.io/xqwb9/ and requests for the program code can be sent via email to DK. Data 198 management, aggregation and plots were performed with MatLab 2022a (The Mathworks, Natick, MA) and inferential statistics with IBM SPSS 27 (IBM, Armonk, NY). DK was supported 199 by grant No. 100019_182146 from the Swiss National Science Foundation. The authors 200 201 declare that they have no conflict of interest. Students participated for class credit and reported normal or corrected-to-normal vision. The study was approved by the ethics 202 committee of the Faculty of Psychology and Educational Sciences at the University of Geneva 203 and was carried out in accordance with the Code of Ethics of the World Medical Association 204 205 (Declaration of Helsinki). Informed written consent was given before the experiment started. Data were collected between 2021 and 2023. 206

Apparatus. The stimuli were displayed on a 22.5-inch LCD monitor (100 Hz, 1,920 ×
 1,200 pixels, standard backlight; VPixx Technologies Inc., Saint-Bruno, Canada). Colors were
 measured with an i1Display Pro (VPixx Edition) colorimeter by X-Rite (Grand Rapids,
 Michigan, United States). Head position was stabilized with a chin/forehead rest at a viewing
 distance of 66 cm. Responses were collected on a RESPONSEPixx Handheld 5-button
 response box (VPixx Technologies Inc., Saint-Bruno, Canada), which had four buttons

--9--

arranged in a diamond shape and one button in the center. The Psychtoolbox (Brainard,

214 1997; Kleiner et al., 2007) controlled stimulus presentation and response collection.

215 Stimuli. There was a placeholder, a cue, and a target display (see Figure 1A). A central fixation cross (0.2° radius, 0.07° linewidth) was shown throughout. The placeholders were 216 217 four outline rings shown to the left, right, above and below the fixation cross. The center-to-218 center distance between the fixation cross and the rings was 3°. To create the outline rings, 219 a large and a slightly smaller circle were drawn with a linewidth of 1 pixel or 0.02°. The radii 220 of the circles were 1.4° and 1.2°. In the cue display, all outline rings were filled. One ring was filled in the cue color while the remaining three rings were filled with the same light gray as 221 the placeholders. In the target display, a letter T tilted by 90° to the left or right was shown 222 223 inside each placeholder. The bars making up the tilted Ts were 1° long and 0.2° thick. One of 224 the Ts was the target color and the remaining Ts were the nontarget color.

Colors are described in xyY with Y in cd/m². The background was dark gray, xyY = (0.312, 0.332, 24.3) while the placeholders and the gray cues were light gray (0.312, 0.332, 48.8). The cue, target and nontarget colors were sampled along an isoluminant color wheel in a CIELAB -space where distances reflect perceived color differences (Fairchild, 2005). In CIELAB-space, the isoluminant color wheel had a lightness of L* = 59 (corresponding to a luminance of 48.8 cd/m²) and a saturation of 64.

The nontarget color was fixed for each participant and the two target colors differed by -30° or 30° of rotation from the nontarget color. The possible nontarget colors were at a rotation of 0°, 45°, 90°, 135°, 180° and 315° on the color wheel. Figure 1B shows colors at rotations from 0° to 360° from left right. The cue colors deviated by -15° or +15° from the target color. The cue was always drawn from the same target-nontarget relation as the subsequent search display. In trial blocks with one target-nontarget relation, there were 32 237 combinations resulting from crossing cue color (-15°, +15°), cue position (left, right, top, 238 bottom), and target position (left, right, top, bottom). The 32 combinations were shown 239 once in each mini-block and there were eight mini-blocks for a total of 256 trials per block. In trial blocks with two target-nontarget relations, the 64 combinations resulting from crossing 240 target-nontarget relation, cue color (-15°, 15°), cue position (left, right, top, bottom), and 241 242 target position (left, right, top, bottom) were shown once in a mini-block. There were four 243 mini-blocks for a total of 256 trials per block. Trial blocks with fixed or mixed target-244 nontarget relations alternated to have two of each. There were four possible block orders. Half of the participants started with a fixed block, and the other half started with a mixed 245 block. For fixed blocks, the order of the two possible target-nontarget relations was 246 247 counterbalanced. The four possible block orders were crossed with the six possible 248 nontarget colors, requiring 24 participants to counterbalance these variables. Every participant worked through 1,024 trials, except for one who terminated the experiment 64 249 trials too early. Because the number of lost trials was small, we decided to keep this dataset. 250 251 Procedure. A trial started with the presentation of the placeholder display for 700 252 ms. Then, the cue display was shown for 50 ms, followed by the placeholder display for 100 253 ms and the target display for 50 ms. The resulting SOA between cue and target was 150 ms. 254 The placeholder display was shown until a key was pressed. Participants responded to a T 255 rotated counterclockwise by a left button press and to a T rotated clockwise by a right button press. They were instructed to respond as rapidly and accurately as possible while 256 257 ignoring the cue display. Trials with choice errors, anticipations and late responses were 258 followed by visual feedback. We considered trials with RTs longer than 1,500 ms as late. Every 64 trials, the percentage of correct responses and the median RTs in the preceding 259 260 trial block were displayed for at least 3,000 ms during a self-terminated pause.

--11--

261 Results

In the analyses of RTs, we successively excluded trials with late responses (0.2%), 262 choice errors (4.2%), and trials with RTs longer than 2.5 SDs above the respective condition 263 mean (2.2%). Cueing effects were calculated by subtracting RTs with valid cues from RTs 264 with invalid cues. We calculated tuning scores by subtracting cueing effects with +15° cue 265 266 colors from cueing effects with -15° cue colors. Positive numbers indicate tuning away from the nontarget color, which is consistent with optimal tuning. Mean absolute RTs, cueing 267 effects and tuning scores are shown in Figures 3 and 4, but our focus was on the analysis of 268 tuning scores. Significant tests remained significant after controlling for false discovery rate 269 (Benjamini & Hochberg, 1995) and we therefore report the uncorrected p-values. 270

We evaluated differences in tuning scores between trial blocks with fixed or mixed target-nontarget relations (see Figure 3). Consistent with an effect of trial history on optimal tuning, tuning scores were larger with fixed than mixed target-nontarget relations (45 vs. 15 ms), t(23) = 4.99, p < .001, Cohen's $d_z = 1.02$. This difference corresponds to the significant three-way interaction in the ANOVA with all relevant factors shown in Table 1. One-sample t-tests showed that both tuning scores were significantly different from zero, ts > 3.49, ps < .002, Cohen's $d_z > 0.71$.

Next, we evaluated effects of intertrial transition in blocks with mixed targetnontarget relations (see Figure 4). We separated trials where the cue matched the previous target-nontarget relation from trials where it did not match. Tuning scores were larger when the cue matched the previous target-nontarget relation than when it did not match (29 vs. 3 ms), t(23) = 2.77 p = .011, Cohen's $d_z = 0.57$. This difference corresponds to the significant three-way interaction in the ANOVA with all relevant factors shown in Table 2. One-sample t-tests showed that tuning scores were significantly different from zero when the cue 285 matched the previous target-nontarget relation (29 ms), t(23) = 4.31, p < .001, Cohen's $d_z =$ 286 0.88, but not when it did not match (3 ms), t(23) = 0.49, p = .628, Cohen's $d_z = 0.10$.

287 We also evaluated whether optimal tuning in fixed blocks was different from trials in mixed blocks where the cues matched the previous target-nontarget relation. In mixed 288 blocks, cues were preceded by as few as one matching target-nontarget relation, whereas 289 290 there were many more repetitions in fixed blocks. Tuning scores tended to be larger in fixed blocks than in matching trials of mixed blocks (42 vs. 29 ms), t(23) = 2.04, p = .053, Cohen's 291 d_z = 0.42. Thus, repetition of the same target-nontarget relation throughout a block of trials 292 tended to increase optimal tuning relative to repetitions occurring randomly in mixed blocks. 293 294 Finally, we performed the same analyses on the proportion of choice errors to rule 295 out speed-accuracy tradeoff. We found no significant differences between tuning scores, ps > .689, but tuning scores were significantly different from zero in fixed blocks (1.9%), t(23) =296 2.71, p = .012, Cohen's $d_z = 0.55$, and approached significance in mixed blocks (1.7%), t(23) =297 2.03, p = .054, Cohen's $d_z = 0.42$. In mixed blocks, tuning scores approached significance on 298 299 matching trials (2.1%), t(23) = 1.91, p = .069, Cohen's $d_z = 0.39$, but not on nonmatching trials (1.4%), p = .225. Because tuning scores in choice errors were not opposite to those in RTs, 300 301 there is no evidence for speed-accuracy tradeoff.

302 Discussion

In Experiment 1, the target-nontarget relations were opposite around the same nontarget color. Tuning scores were larger in fixed than in mixed blocks, and in mixed blocks, tuning scores were larger when the cue color matched the previous target-nontarget relation. These results suggest that trial history played an important role. In fact, tuning scores were not significantly different from zero for cues that did not match the previous target-nontarget relation. Therefore, optimal tuning does not occur simultaneously in

309	opposite directions from the same nontarget color, but only in the direction of the previous
310	target-nontarget relation. In the example with two animals in the grasslands, it would not be
311	possible to tune attention optimally when switching between the tiger and snake, but only
312	when repeatedly searching for the same animal.
313	Experiment 2
314	In Experiment 2, the two target-nontarget relations were separated by at least 90°
315	from each other, resulting in large differences between the colors. In the example from the
316	introduction, this situation corresponds to two different scenes with two different animals.
317	For instance, the observer may switch between looking for a tiger in the grasslands and
318	looking for an almost submerged hippo in a pond. Thus, the target-nontarget relations are
319	independent. One is more orange than the yellow-green grasslands and the other is redder
320	than the brown pond.
321	Method
322	The method was as in Experiment 1 with the following exceptions. The two nontarget
323	colors for each participant were separated by at least 90° to create two independent target-
324	nontarget relations. The five possible pairs of nontarget colors were 0°/90°, 45°/135°,
325	90°/180°, 135°/315°, and 315°/45°. These pairs were crossed with four possible block orders,
326	requiring 20 participants (3 male; age: $M = 21.4$ years, $SD = 3.7$). Data from one participant
327	was replaced because the percentage of errors was excessive compared to the rest of the
328	sample (16% vs. <i>M</i> = 4.8%, <i>SD</i> = 2.6%).
329	Results

Before calculating individual mean RTs per condition, we successively excluded trials with late responses (0.1%), choice errors (4.0%), and trials with RTs longer than 2.5 *SD*s above the respective condition mean (2.0%). Tuning scores were larger with fixed than

--14--

mixed target-nontarget relations (50 vs. 30 ms), t(19) = 3.29, p = .004, Cohen's $d_z = 0.74$. This difference corresponds to the significant three-way interaction in the ANOVA with all relevant factors (see Table 1). One-sample t-tests showed that both tuning scores were significantly different from zero, ts > 5.79, ps < .001, Cohen's $d_z > 1.29$. However, optimal tuning did not differ between trials where the cue matched the previous target-nontarget relation and trials where it did not (35 vs. 26 ms), t(19) = 0.90, p = .382, Cohen's $d_z = 0.20$. The Bayes factor was 4.01 for this comparison, which provides

340 moderate evidence for the null hypothesis. This comparison corresponds to the non-

341 significant three-way interaction in the ANOVA with all relevant factors (see Table 2). One-

342 sample t-tests showed that tuning scores were significantly different from zero when the cue

matched the previous target-nontarget relation (35 ms), t(19) = 4.88, p < .001, Cohen's $d_z =$

1.09, and when it did not (26 ms), t(19) = 4.18, p < .001, Cohen's $d_z = 0.93$.

Further, tuning scores were larger in fixed blocks than in matching trials of mixed blocks (50 vs. 35 ms), t(19) = 2.62, p = .017, Cohen's $d_z = 0.59$, suggesting that repetitions of the same target-nontarget relation throughout a block of trials resulted in stronger optimal tuning than repetitions occurring randomly in mixed blocks.

349 Finally, we performed the same analyses on the proportion of choice errors but 350 found no significant differences between tuning scores, *ps* > .209. However, tuning scores 351 were significantly different from zero in fixed blocks (3.0%), t(19) = 2.78, p = .012, Cohen's d_z = 0.62, and approached significance in mixed blocks (1.5%), t(19) = 2.09, p = .050, Cohen's d_z 352 353 = 0.47. In mixed blocks, tuning scores were significant on matching trials (2.4%), t(19) = 2.19, 354 p = .041, Cohen's $d_z = 0.49$, but not on nonmatching trials (0.5%), p = .726. Because tuning 355 scores in choice errors were not opposite to those in RTs, there is no evidence for speedaccuracy tradeoff. 356

357 Discussion

358	As in Experiment 1, tuning scores were reduced in mixed compared to fixed blocks,
359	but analysis of mixed blocks showed no difference between cues matching or not matching
360	the previous target-nontarget relation. In addition, tuning scores were significant for
361	nonmatching cues, which was not the case in Experiment 1. Thus, attention could be
362	simultaneously tuned to two target-nontarget relations. In the example with two animals in
363	two different scenes, it would be possible to tune attention optimally to the tiger in the
364	grasslands and the hippo in the pond, albeit less optimally than to a single target.
365	Experiments 3
366	In Experiment 3, the target color was fixed and the nontarget colors were opposite.
367	This situation corresponds to the search for a tiger in two different scenes. For instance,
368	there may be one tiger in the yellow-green grasslands and another in an orange-red clay pit.
369	When the two scenes switch randomly, optimal tuning of attention is not warranted because
370	the optimal tuning to one scene may make the target harder to see in the other scene. In the
371	grasslands, attention is optimally tuned to "more orange", but looking for "more orange" is
372	not useful in the orange-red clay pit, where it would be best to look for "yellower".
373	Method
374	The target color was fixed for each participant and the two nontarget colors differed
375	by -30° or 30°. The possible target colors were 0°, 45°, 90°, 135°, 180° and 315°. The four
376	possible block orders were crossed with the six possible target colors, requiring 24
377	participants (3 male; age M = 23.4 years, SD = 7.2).
378	Results

Before calculating individual mean RTs per condition, we successively excluded trials with late responses (0.1%), choice errors (4.7%), and trials with RTs longer than 2.5 *SD*s

--16--

above the respective condition mean (2.0%). Tuning scores were larger in fixed than mixed blocks (31 vs. -6 ms), t(23) = 4.48, p < .001, Cohen's $d_z = 0.91$. This difference corresponds to the significant three-way interaction in the ANOVA with all relevant factors (see Table 1). One-sample t-tests showed that the tuning scores in fixed blocks were significantly different from zero, t(23) = 5.73, p < .001, Cohen's $d_z > 0.71$, but not in mixed blocks (-6 ms), t(23) =1.15, p = .261, Cohen's $d_z = 0.23$.

In mixed blocks, tuning scores were larger on trials where the cue matched the previous target-nontarget relation than when it did not (2 vs. -19 ms), t(23) = 2.94, p = .007, Cohen's $d_z = 0.60$. This difference corresponds to the significant three-way interaction in the ANOVA with all relevant factors (see Table 2). One-sample t-tests showed that the tuning scores with matching cues were not significantly different from zero (2 ms), t(23) = 0.33, p =.742, Cohen's $d_z = 0.07$, whereas they were significantly below zero with nonmatching cues (-19 ms), t(23) = 2.98, p = .007, Cohen's $d_z = 0.61$.

Further, tuning scores were larger in fixed blocks than in matching trials of mixed blocks (31 vs. 2 ms), t(23) = 3.03, p = .006, Cohen's $d_z = 0.62$, suggesting that repetitions of the same target-nontarget relation throughout a block of trials resulted in stronger optimal tuning than repetitions occurring randomly in mixed blocks.

Finally, we performed the same analyses on the proportion of choice errors. Tuning scores tended to be larger in fixed than mixed blocks (2.5% vs. 0.4%), t(23) = 2.00, p = .057, Cohen's $d_z = 0.41$, and tuning scores were significantly different from zero in fixed blocks (2.5%), t(23) = 4.14, p < .001, Cohen's $d_z = 0.85$. None of the remaining differences was significant, ps > .188. Because tuning scores in choice errors were not opposite to those in RTs, speed-accuracy tradeoff is unlikely.

404 Discussion

405 As in the previous experiments, optimal tuning was larger in fixed than in mixed blocks, even compared to those trials in mixed blocks where the cue matched the previous 406 target-nontarget relation. While these results are as in Experiment 1, the magnitude of 407 optimal tuning was smaller. With cues matching the previous target-nontarget relation, 408 409 tuning scores were not significantly different from zero, and with nonmatching cues, tuning scores were significantly below zero. The tuning scores of zero with cues matching the 410 previous target-nontarget relation are surprising because positive tuning scores were 411 observed in these conditions in Experiments 1 and 2. Further, the inverted tuning scores 412 suggest that attention was tuned toward the nontarget color, which is non-adaptive because 413 414 it makes the target more difficult to see.

415

General Discussion

We evaluated effects of trial history in the optimal tuning of attention. Optimal 416 tuning of attention refers to changes in the target representation that decrease the overlap 417 418 between target and nontarget features. That is, participants do not search for the true target color, but for a color that is shifted away from the nontarget color. Previous research has 419 420 established that the memory representation of the target is biased (Chapman et al., 2023; 421 Hamblin-Frohman & Becker, 2021; Jung et al., 2021; Kerzel, 2020; Maith et al., 2021; Scolari 422 & Serences, 2009; Yu & Geng, 2019). Because memory representations are the basis for goal-driven search (Carlisle et al., 2011; Duncan & Humphreys, 1989; Eimer, 2014; Huynh 423 Cong & Kerzel, 2021; Schneider, 2013), the bias may suggest that optimal tuning is a goal-424 425 driven process. However, goal-driven processes may not be the only drivers of optimal tuning. The current study investigated effects of trial history by comparing trial blocks with 426

fixed and mixed target-nontarget relations and by evaluating effects of the match betweenthe current cue color and the previous target color in mixed blocks.

To measure optimal tuning, we used cueing effects. If attention was tuned away from the nontarget color, larger cueing effects are expected for cue colors shifted away from the nontarget color compared with cue colors shifted toward. We refer to the difference between cueing effects as tuning scores. Larger tuning scores reflect stronger optimal tuning. We compared fixed blocks with a single target-nontarget relation to mixed blocks where two target-nontarget relations varied randomly (see Figure 3).

In three experiments, we found that tuning scores were larger in fixed than mixed
blocks. Thus, trial history is important for optimal tuning. The reduced optimal tuning in
mixed blocks was observed regardless of the nature of the target-nontarget relation. That is,
the reduction was found with a fixed nontarget color and two opposite target colors
(Experiment 1), with two target-nontarget relations that were far apart (Experiment 2), and
with a fixed target color and two opposite nontarget colors (Experiment 3).

441 Further, the analysis of intertrial transitions in mixed blocks (see Figure 4) showed 442 that optimal tuning was larger if the cue matched the previous target-nontarget relation, at 443 least in Experiments 1 and 3. With target colors that were far apart (Experiment 2), there was evidence that intertrial transitions did not affect optimal tuning. Nonetheless, the 444 445 overall conclusion of the current study is that trial history contributes substantially to the optimal tuning of attention. It appears that optimal tuning is strongest if the cue matches 446 the previous target-nontarget relation and weaker or even absent if it does not. 447 448 Finally, optimal tuning was larger in fixed blocks than in matching trials of mixed 449 blocks, suggesting that optimal tuning increases with repetition of the same target-

450 nontarget relation. Which process underlies these results?

--19--

451 Underlying Processes

Priming of attentional set. Folk and Remington (2008) observed reduced cueing 452 effects on trials where the cue was different from the previous target. However, this 453 reduction only occurred when participants looked for color singletons, not when they looked 454 for a particular color. Folk and Remington suggested that the previous target would prime an 455 456 attentional control setting for a particular color, but only if participants had not established an attentional set for a particular color to begin with (see also results from dual-target 457 458 search, Irons et al., 2012; Kerzel & Grubert, 2022; Kerzel & Witzel, 2019). Because participants searched for singletons in the current study, it may be that there was priming of 459 optimal tuning. That is, cues matching the target-nontarget relation from the previous trial 460 showed optimal tuning because the respective target-nontarget relation was primed. The 461 problem with this account is that it predicts effects of intertrial transition for all types of 462 target-nontarget relations. However, differences between matching and nonmatching cues 463 were absent in Experiment 2 and there was no optimal tuning with cues matching the 464 previous target-nontarget relation in mixed blocks of Experiment 3. 465 Switch costs. In mixed blocks, the target-nontarget relation switched randomly and 466 467 there were costs associated with these switches. The ANOVAs in Table 2 show significant effects of the previous target-nontarget relation in Experiments 1 and 3. That is, RTs were 468 469 significantly longer when the previous target-nontarget relation was different from the current relation compared to when it was the same. These switch costs amounted to 16 and 470 471 12 ms in Experiments 1 and 3, respectively. It may be that switching to a new target 472 interfered with the optimal tuning of attention. Instead, we have argued that optimal tuning

- 473 was reduced because the cue did not match the previous target-nontarget relation. In fact,
- 474 the two cannot be dissociated because on switch trials, the cue was always nonmatching

--20--

475 with respect to the previous target-nontarget relation. Thus, the question is whether the 476 modestly longer RTs on switch trials may account for the reduced optimal tuning. We think that this is unlikely. First, cueing effects result from shifts of attention preceding the search 477 display and are therefore unrelated to processing of the search display. Second, cueing 478 effects have been shown to increase with longer RTs (Ruthruff et al., 2020), but optimal 479 480 tuning decreased with longer RTs on switch trials. Third, attentional control settings can be rapidly changed without reducing cueing effects (Lien et al., 2010). Fourth, and most 481 importantly, the magnitude of the switch costs does not follow the tuning scores. For 482 instance, switch costs were similar in Experiments 1 and 3, yet the tuning scores were much 483 smaller or inverted in Experiment 3. Nonetheless, there is also evidence supporting the role 484 of switch costs. The absence of switch costs in Experiment 2 was accompanied by the 485 absence of intertrial effects in mixed blocks. However, when taken together, the results do 486 not provide a clear picture regarding the relation between switch costs and optimal tuning. 487 Sensory adaptation. Another explanation of the results is in terms of sensory 488 489 adaptation (see p. 11941 in Scolari & Serences, 2009). Sensory adaptation to color is sufficiently rapid (half-life of less than 25 ms, Rinner & Gegenfurtner, 2000) to allow for 490 491 changes after a single trial. Sensory adaptation may occur to the target and nontarget colors 492 in a search display, and additionally, to intermediate colors because they are not linearly 493 separable (Bauer et al., 1996; D'Zmura, 1991). Consequently, cues shifted from the target toward the nontarget color are in the range of sensory adaptation and their processing may 494 495 be reduced. In contrast, the processing of cue colors shifted away from the nontarget color 496 does not suffer from sensory adaptation. As a result, their ability to capture attention may be larger, resulting in increased cueing effects. Thus, sensory adaptation may explain optimal 497 498 tuning for a single target-nontarget relation, but it also explains some effects of intertrial

--21--

499 transition in mixed blocks. For instance, in Experiment 3, the target color was fixed and only 500 the nontarget color changed. In mixed blocks, optimal tuning was inverted if the cue did not 501 match the previous target-nontarget relation. In contrast to the other accounts, sensory adaptation may provide an explanation. Because the target color was fixed, the same cue 502 colors appeared in both target-nontarget relations, but with different signs (see Figure 2C). 503 504 That is, the cue color shifted toward the nontarget (+) in one target relation was also the cue 505 color shifted away from the nontarget (-) in the other target-nontarget relation. On a switch 506 trial, the adapted cue color from the previous trial (+, e.g. in relation 1) becomes the cue color shifted away (-, e.g. in relation 2) on the current trial. As a result, cueing effects are 507 508 reduced for the cue color shifted away, which is opposite to optimal tuning and results in 509 inverted tuning scores. However, sensory adaptation has difficulty explaining the results 510 from Experiment 1. In Experiment 1, optimal tuning was eliminated when the targetnontarget relation changed. However, sensory adaptation predicts that cue colors from both 511 target-nontarget relations that deviate towards the nontarget color (+15°) should adapt. 512 513 Because adaptation was intermittent, sensory adaptation may be weaker on switch trials, 514 but should nonetheless be present, which was not the case. To obtain more direct evidence 515 for sensory adaptation, measurements of event-related potentials could evaluate sensory 516 adaptation by focusing on an early lateralized potential at posterior electrodes (Kerzel & 517 Huynh Cong, 2021; Schönhammer et al., 2020).

518 Overall, more research is required to clarify which process underlies effects of trial 519 history on optimal tuning. Note that our list of mechanisms may not be exclusive. In the 520 discrimination learning literature, a phenomenon with similar structure has received several 521 interpretations, but is also awaiting a conclusive explanation. When animals are rewarded 522 for responding to a colored light (S+), and not rewarded for responding to another colored

--22--

light (S-), they will respond most to colors on the opposite side of the S+ from the S-, which is
reminiscent of the larger cueing effects for cue colors shifted away from the nontarget
colors. Famously, the learning effects have been attributed to relational responding (Köhler,
1938) or gradients of excitation and inhibition (Spence, 1937). However, neither mechanism
was confirmed in more recent research (Lazareva et al., 2005), calling for more research and
new theories.

529 Conclusions

530 We asked whether trial history contributes to optimal tuning. Optimal tuning occurs when participants search for a target that exaggerates the true relation between target and 531 nontarget colors to improve the signal-to-noise ratio. We observed that optimal tuning was 532 533 stronger in trial blocks with a fixed target-nontarget relation compared with trial blocks 534 where two target-nontarget relations were randomly mixed. Analysis of intertrial transitions in mixed blocks showed that effects of trial history changed with the type of target-535 nontarget relation. When the target-nontarget relations were opposite around the same 536 537 nontarget color (Experiment 1), optimal tuning only occurred when the cue matched the 538 previous target-nontarget, but not when it did not match. This difference was not observed 539 when the target-nontarget relations were far apart (Experiment 2). Finally, when the target-540 nontarget relations were opposite around the same target color (Experiment 3), optimal 541 tuning did not occur with matching cues and was inverted with nonmatching cues. Taken together, these results clearly show that optimal tuning depends on trial history. Several 542 543 mechanisms underlying these effects, such as priming of attentional set, switch costs and

--23--

sensory adaptation, are discussed. However, none of the mechanism can fully account for 544 545 the results. Statement on the generality of findings 546 The participants in the current study were mostly female first-year psychology 547 students, which limits the generalizability of the current findings. 548 549 Acknowledgments 550 Thanks to Christoph Witzel for helping with the CIELAB color space and to Julia 551 Benitez, Matteo Favetta, Inès Lepreux, Alexandre Fortuna Pacheco, Mathieu Zaugg, and Quentin Zongo for helping with data collection. 552 References 553 554 Awh, E., Belopolsky, A. V., & Theeuwes, J. (2012). Top-down versus bottom-up attentional control: a 555 failed theoretical dichotomy. Trends in Cognitive Sciences, 16(8), 437-443. https://doi.org/10.1016/j.tics.2012.06.010 556 Bauer, B., Jolicoeur, P., & Cowan, W. B. (1996). Visual search for colour targets that are or are not 557 linearly separable from distractors. Vision Research, 36(10), 1439-1466. 558 559 https://doi.org/http://dx.doi.org/10.1016/0042-6989(95)00207-3 560 Becker, S. I., Martin, A., & Hamblin-Frohman, Z. (2019). Target templates in singleton search vs. 561 feature-based search modes. Visual Cognition, 27(5-8), 502-517. 562 https://doi.org/10.1080/13506285.2019.1676352 563 Benjamini, Y., & Hochberg, Y. (1995). Controlling the false discovery rate: A practical and powerful approach to multiple testing. Journal of the Royal Statistical Society. Series B 564 565 (Methodological), 57(1), 289-300. Brainard, D. H. (1997). The Psychophysics Toolbox. Spatial Vision, 10, 433-436. 566 567 https://doi.org/10.1163/156856897x00357 568 Carlisle, N. B., Arita, J. T., Pardo, D., & Woodman, G. F. (2011). Attentional templates in visual working 569 memory. The Journal of Neuroscience, 31(25), 9315-9322. https://doi.org/10.1523/jneurosci.1097-11.2011 570 571 Carmel, T., & Lamy, D. (2015). Towards a resolution of the attentional-capture debate. Journal of 572 Experimental Psychology: Human Perception and Performance, 41(6), 1772-1782. 573 https://doi.org/10.1037/xhp0000118 574 Carrasco, M. (2011). Visual attention: The past 25 years. Vision Research, 51(13), 1484-1525. 575 https://doi.org/10.1016/j.visres.2011.04.012 576 Chapman, A. F., Chunharas, C., & Störmer, V. S. (2023). Feature-based attention warps the 577 perception of visual features. Scientific Reports, 13(1), 6487. 578 https://doi.org/10.1038/s41598-023-33488-2 579 D'Zmura, M. (1991). Color in visual search. Vision Research, 31(6), 951-966. 580 https://doi.org/10.1016/0042-6989(91)90203-H 581 Duncan, J., & Humphreys, G. W. (1989). Visual search and stimulus similarity. Psychological Review, 582 96, 433-458. https://doi.org/10.1037/0033-295x.96.3.433

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Table 1. Results from repeated-measures ANOVAs on mean individual reaction times in

682 Experiments 1-3.

683

	Experiment 1		Ex	periment	2	Ex	Experiment 3		
factor	F	p	η_p^2	F	р	η_p^2	F	p	η_p^2
- block	19.10	< .001	.454	10.36	= .005	.353	49.75	< .001	.684
- cue color	3.35	= .080	.127	7.91	= .011	.294	11.10	= .003	.325
- cue validity	109.27	< .001	.826	72.96	< .001	.793	12.30	= .002	.349
- block x cue color		-		3.58	= .074	158	19.02	< .001	.453
 block x cue validity 		-			-		7.16	= .013	.237
- cue color x cue validity	68.91	< .001	.750	67.21	< .001	.780	13.12	= .001	.363
- three-way interaction	24.90	< .001	.520	10.85	= .004	.363	20.07	< .001	.466

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685

686	Note. The factorial design of the ANOVAs was 2 (block: fixed, mixed) × 2 (cue color: -15°, 15°)
687	× 2 (cue validity: valid, invalid). The degrees of freedom were (1, 23) in Experiments 1 and 3,
688	and (1, 19) in Experiment 2. Only significant effects or effects approaching significance (p <
689	.100) are shown for clarity.

691 **Table 2.** Results from repeated-measures ANOVAs on mean individual reaction times in

692 mixed blocks of Experiments 1-3.

693

	Experiment 1		Experiment 2			Ex	Experiment 3		
factor	F	р	η_p^2	F	р	η_p^2	F	р	η_p^2
- previous TNR	13.19	< .001	.364		-		13.77	= .001	.375
- cue color		-			-		23.90	< .001	.510
- cue validity	91.07	< .001	.798	50.25	< .001	.726		-	
- previous TNR x cue color		-			-		7.32	= .013	.241
- cue color x cue validity	13.87	= .001	.376	45.13	< .001	.704		-	
- three-way interaction	7.67	= .011	.250	0.80	= .382	.040	8.64	= .007	.273

694

695

696 **Note.** The factorial design of the ANOVAs was 2 (previous target-nontarget relation: match,

no match) × 2 (cue color: -15°, 15°) × 2 (cue validity: valid, invalid). The degrees of freedom

were (1, 23) in Experiments 1 and 3, and (1, 19) in Experiment 2. Non-significant effects are

699 omitted except for the three-way interaction in Experiment 2. TNR = target-nontarget

700 relation

702 Figure 1.





Note. Panel A illustrates the time course of a trial. Panel B illustrates the selection of cue
colors. In a CIELAB-based color space, the cue colors deviated either away (-15°) or toward
(+15°) the nontarget color. Optimal tuning assumes that attention is tuned away from the
nontarget color. Panel C illustrates the data analysis. Subtracting valid from invalid cue trials
yields cueing effects. Subtracting cueing effects with the +15° cue color from the -15° cue
color yields tuning scores. Positive tuning scores indicate that the cueing effects were larger
for cues deviating away from the nontargets, as predicted by optimal tuning.

713 Figure 2.





715

Note. In each Experiment, there were two target-nontarget relations. In Experiment 1, the
nontarget color was fixed and the target color varied (panel A). In Experiment 2, the targetnontarget relations were far apart (panel B). In Experiment 3, the target color was fixed and
the nontarget color varied (panel C). The signs indicate cue colors deviating away (-) or
towards (+) the nontarget color. t = target.

722 Figure 3.



723 Results in blocks with fixed and mixed target-nontarget relations in Experiment 1-3.

Note. The different dependent variables are shown column by column and the data from the 725 different experiments are shown row by row. In the first and second column, the y-axis 726 shows reaction times in milliseconds (ms). Data from trial blocks with fixed and mixed target-727 728 nontarget relations are shown in the first and second column, respectively. Trials were 729 averaged as a function of cue color $(-15^\circ, +15^\circ)$ and cue validity (iv = invalid, v = valid). The third column shows cueing effects (invalid - valid) in milliseconds on the y-axis. Cueing 730 effects are shown for each cue color in fixed and mixed blocks. The fourth column shows 731 attentional tuning scores in milliseconds on the y-axis. Tuning scores are shown for fixed and 732 mixed blocks. Error bars show the between-subject standard error. Error bars were mostly 733 734 smaller than the symbols. Red asterisks in the rightmost column indicate significant *t*-tests.

735 Figure 4.



736 *Results from Experiment 1-3, data from blocks with mixed target-nontarget relations.*

Note. The different dependent variables are shown column by column and the data from the 738 experiments are shown row by row. In the first and second column, the y-axis shows 739 reaction times in milliseconds (ms). Data from trial blocks where the cue matched the 740 741 previous target-nontarget relation (match) and where it did not match (nonmatch) are 742 shown in the first and second column, respectively. Trials were averaged as a function of cue color (-15°, +15°) and cue validity (iv = invalid, v = valid). The third column shows cueing 743 effects (invalid - valid) in milliseconds on the y-axis. Cueing effects are shown for each cue 744 745 color in matching and nonmatching trials. The fourth column shows attentional tuning scores in milliseconds on the y-axis. Tuning scores are shown for matching and nonmatching 746 trials. Error bars show the between-subject standard error, but were often smaller than the 747

- 748 symbols. Red asterisks in the rightmost column indicate significant t-tests. nonmat. =
- 749 nonmatch